

Objective: Derive optimal dissimilarity criteria by tailoring ranking random processes in the two-sample problem.

Two-sample problem: Let $\{\mathbf{X}_i\}_{i \leq n}$, $\{\mathbf{Y}_j\}_{j \leq m}$ be observations drawn from two samples of unknown probability distribution. By ranking the first sample's data amongst the pooled sample, being able to distinguish possible differences between both distributions.

Contributions

- Express empirical performance criteria as *linear rank statistics* by using Hajek projection and Hoeffding decomposition of U -statistics technics.
- Analyse concentration properties of this novel class of *linear rank processes* when it is generalized with unknown scoring-generating function and optimized over the class of measurable scoring functions.
- In-depth understanding of both global and local dissimilarities criteria for the two-sample problem.
- Apply *linear rank processes* in the two-sample problem and nonparametric homogeneity tests in high dimension.

Notations and Framework

- Let $\mathbf{X} \sim G$, $\mathbf{Y} \sim H$ two independent absolute continuous *r.v.* in the probability space $(\mathcal{X}, \mathcal{P}(\mathcal{X}))$ and consider $\{\mathbf{X}_i\}_{i \leq n}$, $\{\mathbf{Y}_j\}_{j \leq m}$ its realizations *s.t.* p proportion of \mathbf{X} in the pooled sample. Denote by F the *c.d.f.* of the pooled sample *s.t.* : $F := pG + (1-p)H$.
- Let \mathcal{S} by the major class of *scoring functions* *s.t.* $\mathcal{S} := \{s : \mathcal{X} \mapsto \mathbb{R} \text{ measurable}\}$ that maps observations into the real line where its natural relation order can be used. \mathcal{S} has a VC-dimension denoted by \mathcal{V} .
- Once observations are mapped with $s \in \mathcal{S}$, denote by G_s (*resp.* H_s , F_s) the *c.d.f.* of *resp.* $s(\mathbf{X})$ (*resp.* $s(\mathbf{Y})$), $F_s := pG_s + (1-p)H_s$.
- Denote by Ψ the likelihood ratio defined by $\Psi : x \in \mathcal{X} \mapsto \frac{dG(x)}{dH(x)}$.

Related work

- Linear rank statistics* were initially introduced in semi/nonparametric univariate framework by [7], [5].
- Empirical risk minimization of bivariate loss function has been shown to be equivalent with empirical maximization of the R -statistic associated with ([2]).
- Hypothesis testing has been widely studied in univariate and mostly parametric framework.
- Homogeneity testing for the two-sample problem has recently gained interest for multivariate distribution-free settings, especially through the work of Gretton [4] by introducing the Maximum Mean Discrepancy.

Linear rank processes

Definitions: The W -ranking performance measure for two samples is defined by:

$$W_\phi(s) = \mathbb{E}[\phi(F_s(s(\mathbf{X})))], \quad \forall s \in \mathcal{S}. \quad (1)$$

Let $n, m \in \mathbb{N}^*$, the *empirical W -ranking performance measure for two samples* $\{\mathbf{X}_i\}_{i \leq n}$, $\{\mathbf{Y}_j\}_{j \leq m}$ has the following empirical risk functional:

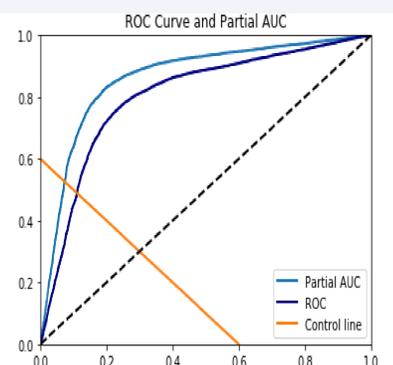
$$\widehat{W}_{n,m}(s) = \sum_{i=1}^n \phi\left(\frac{\text{Rank}(s(\mathbf{X}_i))}{N+1}\right), \quad \forall s \in \mathcal{S}. \quad (2)$$

The function $\phi : [0, 1] \mapsto [0, 1]$ is called the *score-generating function* of the *rank process* $\{\widehat{W}_{n,m}(s)\}_{s \in \mathcal{S}}$. It is supposed to be fixed, nondecreasing and continuously twice differentiable.

Choice of the score-generating function ϕ

Scoring-generating function	Empirical Ranking process	Related Statistic
$\phi = \text{Id}_{[0,1]}$	$\widehat{W}_{n,m}(s) = \frac{1}{N+1} \sum_{i=1}^n \sum_{j=1}^N \mathbb{I}\{s(\mathbf{Z}_j) \leq s(\mathbf{X}_i)\}$	Mann-Whitney-Wilcoxon [1], $\widehat{W}_{n,m}(s) = nm\text{AUC}_{n,m}(s) + \frac{n(n+1)}{2}$
$\phi : u \mapsto u \cdot \mathbb{I}_{\{u \geq u_0\}}$, $u_0 \in (0, 1)$	$\widehat{W}_{n,m}(s) = \frac{1}{N+1} \sum_{i=1}^n \text{Rank}(s(\mathbf{X}_i)) \mathbb{I}\{\text{Rank}(s(\mathbf{X}_i)) \geq u_0(N+1)\}$	Local AUC, concentrates the decision rule on the "best" instances [3]
$\phi : u \mapsto u^q$	$\widehat{W}_{n,m}(s) = \frac{1}{(N+1)^q} \sum_{i=1}^n \text{Rank}(s(\mathbf{X}_i))^q$	Related to q -norm push [6]

Table: Examples of different choices of scoring generating functions



Optimality

Let $n, m \in \mathbb{N}^*$, express:

$$\widehat{W}_{n,m}(s) = \sum_{i=1}^n \phi\left(\frac{N\widehat{F}_{s,N}(s(\mathbf{X}_i))}{N+1}\right), \quad \forall s \in \mathcal{S}. \quad (3)$$

where $\widehat{F}_{s,N}$ is the empirical *c.d.f.* of the scored pooled sample.

A widely used tool for measuring the performance of a scoring function s is the ROC curve defined by:

$$\text{ROC}(s, \cdot) : \alpha \in [0, 1] \mapsto 1 - G_s \circ H_s^{-1}(1 - \alpha) \quad (4)$$

Goal: Interpret the R -processes as optimal unbiased two-sample statistic through the ROC functional curve.

Consider $\mathcal{S}^* = \{s^* = T \circ \Psi \mid T : [0, 1] \rightarrow \mathbb{R} \text{ strictly increasing}\}$.

Proposition: Assume that the score-generating function ϕ is strictly increasing. Then, we have:

$$\forall s \in \mathcal{S}, \quad W_\phi(s) \leq W_\phi(\Psi). \quad (5)$$

Moreover $W_\phi^* \doteq W_\phi(\Psi) = W_\phi(s^*)$ for any $s^* \in \mathcal{S}^*$.

Consequence: The optimal scoring function $s \in \mathcal{S}$ for the homogeneity two-sample problem is the solution of the empirical maximization of the R -process $\{\widehat{W}_{n,m}(s)\}_{s \in \mathcal{S}}$.

Linearization

Proposition: Let $\mathcal{S}_0 \subset \mathcal{S}$ be a VC-major class of functions and suppose ϕ as in definition. Then:

$$\widehat{W}_{n,m}(s) = n\widehat{W}_\phi(s) + \widehat{V}_{n,m}(s) + \mathcal{O}_{\mathbb{P}}(1), \quad \forall s \in \mathcal{S}_0, \quad (6)$$

up to a centering term for the random process, for n, m :

$$\widehat{V}_{n,m}(s) = \sum_{i=1}^n \Phi_s(s(\mathbf{X}_i)) + \sum_{j=1}^m \Phi_s(s(\mathbf{Y}_j)), \quad (7)$$

where $\Phi_s : x \in \mathbb{R} \mapsto p \int_x^{+\infty} \phi'(F_s(u)) dG_s(u)$.

Uniform bound

Theorem: Under the same assumptions, at ϕ fixed, set the empirical W -ranking performance maximizer $\hat{s}_{n,m} = \arg\max_{s \in \mathcal{S}_0} \widehat{W}_{n,m}(s)$. We have, for any $\delta \in (0, 1)$, for $N \gg \max(\kappa_1 + \kappa_2 \log(c/\delta), \sqrt{\kappa_3})$, and with probability at least $1 - \delta$:

$$W_\phi^* - W_\phi(\hat{s}_{n,m}) \leq \kappa \sqrt{\frac{\mathcal{V}(\mathcal{S}_0)}{N}} + \kappa' \sqrt{\frac{\log(2/\delta)}{N}} \quad (8)$$

for some universal positive constants κ, κ', c depending on p , bounds of ϕ and its derivatives, $\mathcal{V}(\mathcal{S}_0)$. $(\kappa_i)_{i \in \{1,2,3\}}$ depend also on δ .

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